

# **Corporate Financing Decisions When Investors Take the Path of Least Resistance**

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We argue that inertial behavior on the part of investors can have significant consequences for corporate financial policy. One implication of investor inertia is that it improves the terms for the acquiring firm in a stock-for-stock merger, since acquirer shares are placed in the hands of investors, who, independent of their beliefs, do not resell these shares on the open market. In the presence of a downward-sloping demand curve, this leads to a reduction in price pressure, and hence to cheaper equity financing. We develop a simple model to illustrate this idea, and present supporting empirical evidence. Both individual and institutional investors tend to hang on to shares granted them in mergers, with this tendency being much stronger for individuals. Consistent with the model and with this cross-sectional pattern in inertia, acquirers targeting firms with high institutional ownership experience more negative announcement effects and greater announcement volume. Moreover, the results are strongest when the overlap in target and acquirer institutional ownership is low and when the demand curve for the acquirer's shares appears to be steep.

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## I. Introduction

Much of finance theory rests on the assumption that investors continuously monitor their portfolios and condition their investment decisions on the most recently available information. Even in models with transaction costs (e.g., Constantinides (1986)) or behavioral biases (e.g., Barberis, Shleifer and Vishny (1998), Daniel, Hirshleifer and Subrahmanyam (1998), and Hong and Stein (1999)), where trade may not be continuous and updating may not be fully rational, investors still can be thought of as processing new information and re-evaluating the decision of whether or not to trade on a constant basis.

While this assumption is convenient for modeling purposes, it is also unrealistic. A large body of existing evidence – which we add to below – suggests that people often behave in a way that might be characterized as inertial, or as taking the path of least resistance. Inertial behavior can arise from a variety of sources, including endowment effects (Thaler (1980), Kahneman, Knetsch and Thaler (1990, 1991)), a tendency to procrastinate in decision making (Akerlof (1991), O’Donoghue and Rabin (1999)), or the cognitive fixed costs associated with re-evaluating and re-optimizing an existing portfolio.<sup>1</sup>

In this paper, we argue that investor inertia may exert a significant influence on financial-market outcomes. Our particular focus is on the consequences of inertia for mergers, and the main idea can be illustrated with a simple example. Consider a firm  $A$  that intends to acquire another firm  $T$  via a stock-for-stock merger, and suppose that the following two conditions hold. First, there is a downward-sloping demand curve for firm  $A$ ’s shares. This downward-sloping demand curve arises not from asymmetric information, but rather from irreducible differences of

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<sup>1</sup> In a famous experiment, Kahneman, Knetsch and Thaler (1990) demonstrate that students who are randomly endowed with a coffee mug subsequently place a much higher reservation price on it than students who are not given a mug, and who are asked to bid on one. They argue that this behavior arises out of loss aversion (Kahneman and Tversky (1979)).

opinion among investors as to the value of  $A$ 's pre-existing assets. Second, and crucially, some investors in the target firm  $T$  are inertial in the following sense: they will not make the active decision to buy shares in  $A$  in, say, a seasoned equity offering (SEO), but if they are granted these shares in a stock-for-stock merger, they will also not make the active decision to sell them.<sup>2</sup>

Under these conditions, an increase in the fraction of inertial investors makes the stock-for-stock merger more attractive to firm  $A$ . Greater target-firm inertia means that more firm- $A$  shares are simply absorbed by the current  $T$  investors, and thus are not ever floated on the open market. With a downward-sloping demand curve for firm- $A$  shares, this implies a smaller negative price impact, which means that firm  $A$  does not have to give up as many new shares in the merger. Said differently, a stock-for-stock merger changes the default setting for inertial  $T$  investors relative to an SEO – it makes the default one in which they are holders of  $A$  shares, which can be thought of as pushing out the overall demand curve for firm- $A$  stock.<sup>3</sup>

After fleshing out our idea with the aid of a simple model in Section II, we examine some of its empirical implications in Section III. We begin by verifying that our premise of investor inertia is relevant in the context of mergers. Using data on both individuals and institutions, we look at investors' propensity to hold on to shares that they are granted in stock-for-stock mergers. We focus on situations in which a given investor in the target owns none of the acquirer before

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<sup>2</sup> These two assumptions place the model in the broader literature of behavioral corporate finance that considers the impact of less-than-fully-rational investors on corporate investment and financing decisions. Early work in this area includes Fischer and Merton (1984), Morck, Shleifer, and Vishny (1990a), Blanchard, Rhee, and Summers (1993), and Stein (1996). Behavioral papers on the specific topic of mergers include Roll (1986), Shleifer and Vishny (2003), Malmendier and Tate (2003), and Friedman (2004). See Baker, Ruback, and Wurgler (2004) for a survey.

<sup>3</sup> Madrian and Shea (2001), and Choi, Laibson, Madrian and Metrick (2002, 2003) demonstrate just how powerful the effect of defaults can be in the context of retirement savings decisions. To take just one example, when firms set the default in their 401(k) plans to automatic enrollment, very few workers choose to opt out, resulting in participation rates close to 100 percent. In contrast, if the default is no enrollment, so that a worker has to make an active decision to participate in the plan, participation rates are generally much lower. In a related corporate finance paper, Zhang (2004) argues that the endowment effect can explain IPO underpricing.

the deal, so that it can be inferred that he does not have a high valuation for the acquirer.<sup>4</sup> Even in such cases, target investors have a remarkably high likelihood of owning acquirer shares after the merger transaction closes. We estimate that roughly 80 percent of individuals behave as “sleepers”, and simply accept shares they are given in a merger. For institutions, the estimated fraction of sleepers is significantly smaller, at around 30 percent, but still noteworthy.

Next, we test one of the theory’s central predictions. Given that institutional investors are less prone to inertia than individuals, our model implies that the announcement return to the *acquirer* in a stock-for-stock deal will be more negative, all else equal, if the *target* firm has a higher proportion of institutional shareholders. This is because institutions are expected to dump more of the acquirer-firm shares they receive back onto the market; individuals, by contrast, tend to hang on to these shares, thereby mitigating price impact. Using a variety of specifications, we find robust evidence for this hypothesis. We also provide another clue that these return effects are due to price pressure, as our model suggests: the acquirer has more trading volume around the announcement date when the target has a higher proportion of institutional shareholders.

To rule out alternative explanations, we verify some finer predictions of the model. Both acquirer return and volume effects are largest when the overlap between target and acquirer institutional ownership is small. Intuitively, non-overlapping institutional owners of the target are the ones most likely to unload their shares on announcement of a merger, as they have demonstrated a lack of interest in holding acquirer assets. The results are also stronger when various proxies suggest that the acquirer’s demand curve is steep. Finally, we show that, consistent with our model, each of the above results is only present in stock-swap mergers, and not in cash deals.

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<sup>4</sup> As discussed below, the conceptually cleanest case is one in which the acquirer is very large relative to the target, so the post-merger combined company is composed almost entirely of acquirer-firm assets.

On a more speculative note, we make two empirical connections in Section IV between our theory and corporate financing decisions. The first has to do with the means of payment in a merger. In a more general version of the model, acquiring firms prefer to use stock as opposed to cash as consideration when the inertia of target shareholders is high. Consistent with this prediction, we document a negative relationship between target institutional ownership and the probability that a merger is conducted with stock. Like our other results, this effect is most pronounced when the overlap between target and acquirer institutional ownership is small.<sup>5</sup>

A second connection is with recent empirical work by Fama and French (2004). They document that although SEOs are relatively rare, *total* external equity financing – which in addition to SEOs, can come in the form of stock-for-stock mergers or stock-based employee compensation – is actually quite substantial for most classes of firms. For example, over the period from 1999 through 2001, Fama and French find that S&P 100 firms raised an annual average of only 0.09 percent of assets via SEOs, but 1.05 percent via various forms of stock-based compensation, and 3.68 percent via mergers. In other words, the volume of equity finance raised in mergers by these large firms was roughly 40 times that raised in SEOs.

As Fama and French (2004) point out, these stylized facts are not easily reconciled with standard corporate finance theories, such as the asymmetric-information-based approach of Myers and Majluf (1984). Myers and Majluf have a good story for the relative scarcity of SEOs taken in isolation, but they have little to say about why mergers would be a dominant substitute.

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<sup>5</sup> This result is also consistent with institutions having an innate preference for cash. However, because acquirer shares can be converted to cash in the secondary market, a strong preference suggests that institutions expect significant transaction costs. Our theory emphasizes the price pressure component of these costs.

Indeed, as noted above, a direct application of Myers-Majluf logic would seem to imply that stock-for-stock mergers face the same asymmetric-information problems as SEOs.<sup>6</sup>

In contrast, our theory suggests an affirmative rationale for the primacy of stock-for-stock mergers as compared to SEOs.<sup>7</sup> To be specific, imagine a firm with an exogenously specified growth strategy that, over the next year, will involve one acquisition, and one major new greenfield investment (e.g., construction of a new plant). Suppose further that, due to optimal-capital-structure considerations, one of these two transactions will need to be financed with an equity issue – i.e., either the merger will have to be stock-for-stock, or the greenfield investment will have to be accompanied by an SEO. Which outcome is more likely? Since the SEO effectively amounts to a limiting case of our model with no investor inertia, it will be associated with a more negative price impact, all else equal, and hence will tend to be less attractive. Thus we would expect the firm to finance the merger with a stock swap, but to finance the greenfield investment with cash. This is an outcome very much in the spirit of Fama and French (2004).

## II. The Model

### A. *Investor Beliefs*

The model has three dates, labeled 0, 1 and 2. The focus is on the behavior of a potential acquirer firm *A*, which will be faced with an investment decision at time 1. As of time 0,

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<sup>6</sup> It is possible to augment the Myers-Majluf model so as to introduce a distinction between mergers and SEOs. Perhaps the terms of a merger have to be negotiated with target management, who are better informed about the value of acquirer assets than the typical outside shareholder, and who thus act as a certifier of acquirer-firm value. The problem with this story is that it may well be possible for the acquirer to co-opt target management by offering a side payment (e.g., a seat on the board of the combined company) which would compromise their role as a certifier.

<sup>7</sup> Our basic line of reasoning suggests that stock-based employee compensation may also be preferred to SEOs. If workers are subject to an endowment effect (so that once granted stock, they are reluctant to sell it, even if they would not have gone out and bought it on their own in the first place) a firm facing a downward-sloping demand curve will prefer to place stock with them than to sell it on the open market. This observation may help to resolve the puzzle of why firms give stock to low-level employees, where incentive effects are likely to be minimal. See, e.g., Bergman and Jenter (2003) and Oyer (2004).

however, the prospect of investment is unanticipated by the market, so  $A$ 's stock is priced solely on the basis of cash flows from assets already in place. These assets in place will yield a liquidating dividend of  $D$  at time 2.

Our first crucial assumption is that there are differences of opinion among investors in firm  $A$  as to the expected value of  $D$ . In particular, there is a continuum of “ $A$ -specialists” who have values of  $E(D)$  uniformly distributed on the interval  $[F, F + H]$ , where the parameter  $H$  can be interpreted as a measure of the divergence of opinion. And while they are risk-neutral, the  $A$ -specialists are constrained to not invest more than their total wealth of  $W$ . To ensure the existence of interior solutions in what follows, we stipulate that  $W > F$ . Finally, maintaining a short position over the interval from time 0 to time 2 is assumed to be impossible.<sup>8</sup>

Taken together, these assumptions have the effect of creating a downward-sloping demand curve for firm- $A$  assets.<sup>9</sup> Given this demand curve, the market value of the firm at time 0,  $P_0$ , is determined by setting  $P_0$  equal to the total wealth of those  $A$ -specialists with valuations in excess of  $P_0$ . In other words, the value of the firm is equal to the wealth of those investors who are buyers in equilibrium. This condition is equivalent to:

$$P_0 = \frac{W}{H}(F + H - P_0), \text{ or } P_0 = (F + H) \frac{W}{W+H}. \quad (1)$$

From (1), along with our assumption that  $W > F$ , it follows that  $P_0$  always lies between  $F$  and  $(F + H)$ . The fraction of investors who are long the stock in equilibrium is given by  $\frac{F+H}{W+H}$ .

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<sup>8</sup> Miller (1977) was the first to model the combined effects of differences of opinion and short-sales constraints. Recent treatments include Chen, Hong and Stein (2002), Scheinkman and Xiong (2003), and Hong, Scheinkman and Xiong (2004).

<sup>9</sup> We use this formulation to ease the exposition. In fact, any set of assumptions that produces downward-sloping demand is sufficient. And, on this point, the empirical literature is fairly clear. Bagwell (1992) and Hodrick (1999) show that investors have downward-sloping demand in Dutch auction share repurchases. Shleifer (1986), Harris and Gurel (1986) and more recently Kaul, Mehrotra, and Morck (2000), Wurgler and Zhuravskaya (2002), and Greenwood (2004) illustrate the price impact of uninformed demand, by examining index inclusion and rebalancing decisions. Mitchell, Pulvino and Stafford (2004) focus on price pressure in the context of mergers and acquisitions.

Also, we have the intuitive properties that  $\frac{dP_0}{dW} > 0$ , and  $\frac{dP_0}{dH} > 0$ . The latter is just a version of Miller's (1977) insight that, in the presence of a short-sales constraint, prices are increasing in the heterogeneity of investor opinion. To see the import of the downward-sloping demand curve, observe that  $\frac{dP_0}{dF} = \frac{W}{W+H} < 1$ . This means that, if  $W$  is held constant, the firm's market value does not go up one-for-one with an increase in expected cashflows. The intuition is that as the firm gets larger, shares must be absorbed by investors who are less optimistic. Moreover,  $\frac{d^2 P_0}{dF dH} < 0$ , so that an increase in heterogeneity amplifies the downward slope of the demand curve.

**Example:** Suppose  $F = 100$ ,  $H = 100$ , and  $W = 300$ . Then firm  $A$  has a market value of  $P_0 = 150$  at time 0. The more optimistic half of the  $A$ -specialists (those with valuations between 150 and 200) own all the shares, while the remaining half of the  $A$ -specialists (those with valuations between 100 and 150) sit out of the market.

### *B. Stock-for-Stock Merger When All Target Shareholders Are Awake*

At time 1, the manager of firm  $A$  announces that he has decided to acquire a target firm  $T$  and to finance the acquisition with an equity issue. The purchase price of the target is  $K$ , and the merger will increase firm  $A$ 's terminal dividend by  $R > K$ , which implies that the merger is positive-NPV for firm  $A$ .<sup>10</sup> For simplicity, we assume that there is no disagreement among the  $A$ -specialists as to the value added by the acquisition, so that once it is on the books, their expectations of terminal cash flow are uniformly distributed on the interval  $[F + R, F + R + H]$ .

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<sup>10</sup> Although we assume that the  $A$ -manager is interested in the merger simply because it represents a positive-NPV investment, our results would be similar if instead we assumed that the  $A$ -manager was motivated by either: i) a desire to increase the size of his empire; or ii) a belief that his stock was overvalued, as in Stein (1996) or Shleifer and Vishny (2003), for example.



We assume that none of the target shareholders are among the group of *A*-specialists. That is, as of time 0, their expectations of firm *A*'s terminal dividend are relatively low; without loss of generality one can think of all of them as simply having  $E(D) = F$ . Empirically, this implies that, prior to the merger announcement at time 1, there is no overlap between the investors in firms *A* and *T*.<sup>11</sup>

In a stock-for-stock merger, shares in the newly-created merged firm *M* are placed directly into the hands of firm *T*'s investors. Based on their pessimism about the prospects for firm *A*'s pre-existing assets, these *T*-investors also have low valuations of *M* – they consider it to be worth only  $(F + R)$  – as compared to the pool of *A*-specialists, whose valuations of *M* all exceed  $(F + R)$ . Thus one might expect that the *T*-investors would immediately take any shares in *M* that they receive in a stock-for-stock exchange, and sell them off in the open market, where these shares would have to be absorbed by *A*-specialists.

This logic is correct, if all *T*-investors are aware of the merger, and react to it in a rational fashion. Indeed, under full rationality, the stock-for-stock merger is identical to the following “decoupled” transaction: shares in the merged firm *M* are sold off in an SEO (and bought by *A*-specialists) and the proceeds of this SEO are used to pay the *T*-investors in cash for their shares. In either case, the absence of inertia implies that any shares in *M* must ultimately wind up in the hands of those with the highest valuations for *M*, and the *T*-investors are not among this group.

To finance the merger, enough new shares must be issued to raise an amount  $K$ . For the case where all *T*-investors are rational, denote the post-merger market value of the firm at time 1 by  $P_1^R$ . Note that any *A*-specialists who were long at time 0 cannot participate in the new issue, since they already have all of their wealth invested in firm-*A* shares. Thus the new shares must

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<sup>11</sup> Nothing substantive changes if we allow for partial overlap.

be absorbed by those  $A$ -specialists who were previously on the sidelines. This group has total wealth of  $(W - P_0)$ , and has valuations distributed uniformly on the interval  $[F + R, P_0 + R]$ .<sup>12</sup>

Since the market value of the shares they absorb must equal  $K$ , equilibrium requires that

$$K = \frac{W - P_0}{P_0 - F} (P_0 + R - P_1^R), \quad (2)$$

which can be re-written as

$$P_1^R = P_0 + R - \frac{H}{W} K. \quad (3)$$

Note that  $P_1^R$  is necessarily less than  $P_0 + R$  – the market value does not go up by the full amount of the added cashflows  $R$  from the firm- $T$  assets. In other words, the equity issue has a price-pressure effect, the magnitude of which is increasing in the heterogeneity parameter  $H$ . This reflects the fact that the issue must be absorbed by the relatively less optimistic  $A$ -specialists, who were sitting out of the market prior to the issue.

**Example (continued):** Keep  $F = 100$ ,  $H = 100$ , and  $W = 300$ . Suppose further that  $K = 100$ , and  $R = 110$ . This yields  $P_1^R = 226.67$ , so that a 44.12 percent share in the firm is issued in the merger (since  $0.4412 \cdot 226.67 = 100$ ). The market value of the stake held by pre-existing firm- $A$  shareholders drops from 150 to 126.67 – the merger is accompanied by a negative price impact of 15.56 percent. After the merger, 83.33 percent of the  $A$ -specialists have long positions.

### *C. Stock-for-Stock Merger When Some Target Shareholders Are Asleep*

In order to explore the consequences of investor inertia, we now posit that only a fraction  $\alpha < 1$  of the  $T$ -investors are “awake”. As before, these awake  $T$ -investors rationally sell off any

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<sup>12</sup> In order for the equity issue to be able to raise  $K$  from among the pool of  $A$ -specialists, it must be that the total wealth of the previously sidelined  $A$ -specialists,  $(W - P_0)$ , weakly exceeds  $K$ . This condition can be expressed as  $W(W - F)/(W + H) \geq K$ , and we assume it holds in what follows.

shares in the merged firm  $M$  that they receive in the merger. The remaining  $(1 - \alpha)$  of the  $T$ -investors are “asleep”, and simply hold on to the shares in  $M$  that they are given. Thus for the sleeping  $T$ -investors, the default condition matters: they would not actively seek to buy shares offered in an SEO, but they also do not actively seek to sell shares that are granted to them as part of a stock-for-stock merger. Or said differently, the sleeping  $T$ -investors always take the path of least resistance, which is simply to do nothing.

In the presence of sleepers, we assume that the awake  $T$ -investors are the only ones who actively evaluate the bid from firm  $A$ .<sup>13</sup> In doing so, they continue to place a reservation value of  $K$  on their firm’s assets, and recognize that they will immediately resell all shares that they receive in the stock swap. If we denote by  $P_1^S$  the post-announcement market value of the merged firm when there are sleepers, then the requirement that the bid be satisfactory to the awake  $T$ -investors amounts to saying that the total value of shares issued in the merger, evaluated at  $P_1^S$ , be equal to  $K$ . Note that this condition is exactly the same as the analogous one in the fully rational case.

Where things differ is in the determination of  $P_1^S$ . As a result of the sleeping  $T$ -investors who hang on to their shares, only a fraction  $\alpha$  of the shares issued in the merger ever comes on the market, and hence only this fraction must be absorbed by the pool of previously-sidelined  $A$ -specialists. Thus equations (2) and (3) are modified as follows:

$$\alpha K = \frac{W-P_0}{P_0-F} (P_0 + R - P_1^S), \quad (4)$$

$$P_1^S = P_0 + R - \frac{H}{W} \alpha K. \quad (5)$$

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<sup>13</sup> One might imagine that only awake  $T$ -investors are ever involved in negotiating the terms of the merger with the acquirer. A friendly merger typically requires approval of a majority of shareholders, which at first glance would seem to imply a lower bound on  $\alpha$  of 0.50, to the extent that one needs to be awake to vote. However, to the extent that shares of individual investors are held (and voted) in street name, it is possible that  $\alpha$  might fall below 0.50 without preventing a deal from getting done.

**Example (continued):** As before, keep  $F = 100$ ,  $H = 100$ ,  $W = 300$ ,  $K = 100$ , and  $R = 110$ . Assume that the fraction of sleepers among  $T$ -investors is given by  $\alpha = 0.50$ . These parameters yield  $P_I^S = 243.31$ , which implies that a 41.10 percent share of the firm is issued in the merger, less than the 44.12 percent share issued in the case with fully rational investors. (Note that the value of the merger bid, evaluated at market prices, is  $0.411 \cdot 243.31 = 100$ .) The market value of the stake held by pre-existing firm-A shareholders drops from 150 to 143.31 – the merger is accompanied by a negative price impact of 4.46 percent. After the merger, 66.67 percent of the A-specialists have long positions with an aggregate market value of 193.31, and sleeping  $T$ -investors have a long position with a market value of 50.

#### *D. Empirical Implications*

The model's most direct empirical implications can be summarized as follows.

#### **Proposition 1: Merger Announcement Effects**

(i) All else equal, an increase in the fraction  $\alpha$  of awake target-firm investors strengthens the adverse impact of a stock-for-stock merger announcement on the price of the bidding firm. (ii) An increase in the slope of the acquirer's demand curve – as measured by the degree of investor heterogeneity  $H$  – has a similar effect. (iii) There is an interaction between these two variables: a steeper acquirer demand curve amplifies the negative stock-price consequences of awake target-firm investors. Thus, denoting the announcement price impact by  $\Delta P$ , we have that: (i)  $\frac{d\Delta P}{d\alpha} < 0$ ; (ii)  $\frac{d\Delta P}{dH} < 0$ ; and (iii)  $\frac{d^2\Delta P}{d\alpha dH} < 0$ .

Proposition 1 follows immediately from inspection of equation (5). Part (i) of the proposition forms the basis for one of our main empirical tests below. Part (ii) – which does not involve the wakefulness parameter  $\alpha$ , and which holds even in a world with no inertia – is the subject of a recent paper by Moeller, Schlingemann and Stulz (2004). They use the dispersion of analyst forecasts as a proxy for  $H$ , and find evidence consistent with the hypothesis that  $\frac{d\Delta P}{dH} < 0$ . Taking a similar approach to measuring  $H$ , we also attempt to test part (iii) of the proposition. Thus, consistent with our theoretical emphasis, our empirical work centers on those effects that are most directly related to the wakefulness parameter  $\alpha$ .

#### *E. Further Issues*

##### **1. Merger arbitrage**

In the model, the announcement and completion of a merger occur simultaneously, at time 1. More realistically, there can be a substantial time lag between announcement and completion, and completion may not be a sure thing when the deal is first announced. It is easy to extend the model to incorporate these features. Neither of the two propositions above is changed, though the model may now also admit a role for merger arbitrageurs.

To see how such arbitrage might work, imagine that the announcement of a merger occurs at time 1, but the transaction is not completed until time  $1\frac{1}{2}$ . Moreover, since completion is not ensured as of time 1, the stocks of firm  $T$  and firm  $A$  are not interchangeable immediately post-announcement – they are no longer certain to both turn into claims on the merged firm  $M$ . In this setting, it is possible that awake  $T$ -investors will want to sell their shares in  $T$  immediately, as of time 1. But the previously-sidelined  $A$ -specialists may prefer to buy  $A$  shares when their price falls at this time, rather than buying  $T$  shares. This is because their primitive

preference is for  $A$  assets, and there is a risk that the  $T$  shares will not turn into a claim on any  $A$  assets, if the deal falls through. The arbitrageurs can bridge this gap by buying  $T$  shares from the  $T$ -investors, and for each share bought, short-selling  $K/(P_1^S - K)$   $A$  shares to the  $A$ -specialists.<sup>14</sup> If the deal does indeed go through, each side of the arbitrage trade will convert into the same number of shares in the merged firm  $M$ . If not, arbitrageurs will be left with an unhedged position in  $T$  and  $A$  shares.

As long as completion risk is small, the results for prices at time 1 remain approximately the same as described above. That is, the adverse impact to  $A$ 's stock price will occur primarily on announcement of the deal, rather than on completion, and will continue to be a function of the number of sleepers among the  $T$ -investors. Intuitively, the more sleepers there are, the fewer shares of  $T$  are unloaded onto the arbitrageurs at time 1, and hence the fewer shares of  $A$  are short-sold by these arbitrageurs into the downward-sloping demand curve of the  $A$ -specialists.

The one thing we gain by explicitly considering the process of merger arbitrage is a more precise set of predictions about trading volume as of the announcement date.

### **Proposition 2: Abnormal Trading Volume Around Merger Announcements**

(i) All else equal, an increase in the fraction  $\alpha$  of awake target-firm investors leads both to more trading volume in the target around the announcement date; and (ii) to more trading volume in the acquirer.

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<sup>14</sup> This calculation assumes that both firms have the same number of shares outstanding. Note that we are now allowing arbitrageurs to hold short positions between time 1 and time 1½, even though we previously ruled out short-selling by investors over the longer interval from time 1 to time 2. A loose rationalization might be that those players who have the ability to short-sell are unwilling to take long-horizon unhedged short positions because of the fundamental risk involved.

Part (i) of the proposition is self-evident: if all target-firm shareholders are asleep, none of them will sell on announcement of the deal, and there will be no trading volume in the target. Part (ii) is a bit more subtle, and relies on the merger-arbitrage mechanism: the more shares are dumped by awake target-firm shareholders, the more the arbitrageurs have to step in and buy, and hence the more short-selling of the acquirer they end up doing to hedge their positions.<sup>15</sup> These predictions for volume are a useful complement to our predictions for acquirer stock returns in Proposition 1 above. If both Propositions 1 and 2 are borne out in the data, it becomes more likely that the results for stock returns are driven by the sort of price-pressure effects envisioned in our model, as opposed to some other confounding factor.

## **2. Capital gains taxes**

We have not yet addressed the following important question: to what extent can what we call investor inertia, or sleepiness, be thought of as simply a rational reluctance on the part of target-firm shareholders to incur capital-gains taxes by selling their shares? In other words, might we be able to generate our results by appealing to nothing more than such taxes?

The answer is no. On the one hand, it is possible that cross-sectional differences in the degree of sleepiness across target firms could be related to tax considerations.<sup>16</sup> For example, if institutional investors are always tax exempt, and individuals are not, this might explain why individuals are more prone to hang on to their shares after a merger than institutions. But taxes

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<sup>15</sup> An important caveat is that this short-selling will only occur if the terms of the merger involve an exchange ratio that is fixed as of the announcement date. If, instead, the dollar value of the bid is fixed, and the exchange ratio is left to float until completion, the arbitrageurs will not want to short the acquirer. See Mitchell, Pulvino and Stafford (2004) for an analysis of this issue.

<sup>16</sup> This does not seem to be the case empirically, however. We show below that the degree of sleepiness does not seem to vary much with proxies for capital-gains-tax exposure. For example, sleepiness is not noticeably more pronounced in target firms that have had large price run-ups over the past few years. Nor is it any greater among individual investors with taxable as opposed to non-taxable accounts.

alone cannot generate the sorts of patterns in acquirer announcement returns predicted by Proposition 1.

To see why, consider a target firm in which all investors are heavily-taxed individuals. It may be true that these individuals will be reluctant to sell any shares in the acquirer that they get as part of a merger. But if they are fully rational, and understand this *ex ante*, they will demand a greater number of shares in the first place, to compensate them for the fact that they are going to be forced to hold stock that they do not value highly. So on net, the acquirer is made *worse* off as a result of the tax-related friction. As long as everybody is rational, the acquirer must bear the cost of inefficiently placing shares in the wrong hands. In contrast, if target shareholders are non-taxable institutions, the shares will ultimately make their way into the right hands, those of the previously-sidelined *A*-investors who value them more highly.

**Example (continued):** As before, keep  $F = 100$ ,  $H = 100$ ,  $W = 300$ ,  $K = 100$ , and  $R = 110$ . Consider two cases. In the first, all target shareholders are awake, and there are no taxes. As we have seen, this implies that 44.12 percent of the firm's shares must be issued to raise 100. In the second case, all target shareholders are again awake, but there are prohibitive capital-gains taxes, so that target shareholders are forced to hang on to any acquirer shares they receive in a merger. Target shareholders value the combined post-merger firm at 210, and have a reservation price of 100. So to get them to sell, they must be given 47.62 percent of the combined firm ( $100/210 = 0.4762$ ), which is more than in the case with no taxes.

The bottom line is that investor inertia delivers something that capital-gains taxes alone cannot. When it is granted to them as consideration in a merger, not only are inertial target



investors more likely to hang on to a stock that they would otherwise never have bought, but crucially, *they do not have to be compensated for doing so*. This latter feature is why acquirers can benefit from target-investor inertia, as in Proposition 1.

### **III. Empirical Analysis**

Our empirical work is divided into two parts. First, we simply document that the key premise of our model holds in the data: that is, target investors behave in an inertial fashion around merger transactions. While we find that both individual and institutional investors exhibit inertia, it is substantially more pronounced among individuals. Second, using this distinction between individuals and institutions to create a proxy for the extent of investor inertia, we test the cross-sectional implications of the model summarized in Propositions 1 and 2. We show that acquirer announcement returns are more negative in transactions where the target firm has a greater proportion of institutional (i.e., awake) shareholders, and that acquirer abnormal trading volume is also greater around such transactions.

#### *A. Data*

Our sample of mergers includes 3,003 successfully completed transactions announced between the second quarter of 1980 and the fourth quarter of 2000. Of these, most of our analysis will focus on the subset of 1,858 stock-swap deals; the remaining cash deals are used only as a control sample. We require that each deal involves a public acquirer, with a matched announcement return available from CRSP. Quarterly observations on institutional ownership come from the CDA Spectrum Institutional Holdings database, over the period from the first quarter of 1980 through the fourth quarter of 2002. We track institutional holdings starting in the

quarter prior to announcement. Information on individual investor holdings comes from the records of a discount brokerage firm (this is the Odean (1998) data) and is available only from 1991 to 1996. We also make use of a variety of CRSP and Compustat variables, as well as analyst forecast data from IBES.

Table 1 presents summary statistics for the mergers in our sample. All variables are Winsorized at the first and 99<sup>th</sup> percentiles, both in Table 1 and in the analysis that follows. In Panel A, we look at target institutional ownership, which is calculated for the quarter prior to the announcement of a merger and expressed as a percentage of shares outstanding. We also look at “non-overlapping” target institutional ownership, defined as the fraction of the target owned by those institutions who own no shares in the acquiring firm.<sup>17</sup> As we discuss in more detail below, non-overlapping institutional ownership is probably the best available proxy for the wakefulness parameter  $\alpha$ .

Panel B gives several deal characteristics. Acquirer and target size are equal to price times shares outstanding two days prior to announcement. Relative size is equal to target size expressed as a percentage of total target and acquirer size. Acquirer and target leverage are equal to interest bearing debt (items 9+34 from Compustat) expressed as a percentage of book assets (item 6) at the fiscal year end prior to announcement. Acquirer and target market-to-book are equal to book assets minus book equity (items 216-130+35) plus price times shares outstanding (from CRSP), all divided by book assets. Same Industry is an indicator variable equal to one if the target and the acquirer are in the same Fama and French (1997) industry.

Panel C presents stock-market data, all taken from CRSP. Acquirer and target announcement returns are cumulative returns in excess of the value-weighted market over a five-

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<sup>17</sup> Matvos and Ostrovsky (2004) highlight the importance of overlapping institutional ownership in mergers, and discuss its implications for corporate governance.

day window surrounding the announcement of a merger. Acquirer announcement volume is the average daily volume over a five-day window surrounding the announcement of a merger, expressed as a percentage of shares outstanding. Normal volume is the average daily volume over a 60-day window starting 90 trading days before the announcement of a merger.

Finally, Panel D shows two acquirer demand-curve proxies, constructed from data in CRSP and IBES. Our first proxy follows Scherbina (2002), and Moeller, Schlingemann, and Stulz (2004), who use dispersion in analyst forecasts as a measure of disagreement about fundamental value. Our particular measure is the same adopted by Moeller et al in a very similar context, and is equal to the standard deviation of all outstanding earnings forecasts of long-term growth. Our second proxy, idiosyncratic risk, is the standard deviation of the residuals from a regression of acquirer excess returns on the Fama-French factors ( $R_M$ ,  $SMB$ ,  $HML$ ), and the matched 48-industry portfolio return.<sup>18</sup>

## *B. Investor Inertia*

### **1. Methodology**

If all mergers were announced and completed instantaneously – as in our simple model – it would be a straightforward matter to measure target investor inertia. Consider a merger in which we have a set of target investors *who have no initial position in the acquirer*. By revealed preference, these “target-only investors” do not place an especially high value on the acquirer. If these investors were all awake, we would expect them to sell their shares immediately upon announcement and simultaneous completion. In contrast, if the target-only investors were asleep, we would expect them to do nothing. Thus a natural measure of inertia would be the

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<sup>18</sup> All factor and portfolio returns were obtained from Ken French’s website.

fraction of target-only investors doing nothing. This would correspond exactly to the key variable  $(1 - \alpha)$  in our model.

In practice, there is a lag between the announcement and completion of a deal. Our aim is to get a picture of the total selling activity by target-only investors over this interval. Moreover, the passage of time raises subtle benchmarking issues. For example, suppose that a particular deal takes a year to close, and that we see ten percent of the target-only investors selling their shares during the course of this year. Should we draw the conclusion that ten percent are effectively awake, and responding to the merger? Not necessarily. Even in a year without a merger, there is a baseline level of turnover. In other words, we expect to see some selling, for example because of liquidity demand, even if all target investors are completely oblivious to the fact that a merger has occurred. So, to measure inertia correctly, we need to calculate turnover above and beyond what would be expected absent a merger.

Figure 1 illustrates our method for calculating investor inertia in a hypothetical merger transaction. Time from announcement to completion is measured along the X-axis. The merger is announced at the Y-axis and is completed by the end of the dashed line. At any point in between, the dashed line represents the fraction of those pre-announcement target investors with no initial position in the acquirer who continue to hold a long position, either in the target (in the period prior to completion), or in the acquirer (in the period after completion). For example, the figure shows that at completion, 32 percent of the original target-only investors are still holding their positions, which have now converted into shares of the acquirer.

The upper solid line in Figure 1, which starts out at 100 percent and declines gradually, is the “fully passive” benchmark. This is the fraction of target-only investors we would expect to be continuing holders based on normal turnover during a non-merger period. In this hypothetical

example, the figure shows that only 53 percent of investors are expected to still hold their target shares seven quarters later, even if they pay no attention to the merger. The lower solid line, which starts out at zero percent and increases gradually, is the “fully active” benchmark. The idea here is that an investor who is awake and sells the target immediately upon announcement might, over time, experience a change in view, and re-buy the target or the acquirer at some later date. The figure shows that 14 percent of the original target-only investors in this example would be expected to own the acquirer seven quarters later if they were to sell out at the time of the merger announcement, simply because they revise their opinion of the acquirer.<sup>19</sup>

To compute an adjusted measure of inertia that incorporates both of these benchmarks, we define inertia at completion as the difference between the fraction of post-completion holders and the fully active benchmark (labeled  $a$  in Figure 1), divided by the difference between the fully passive benchmark and the fully active benchmark (labeled  $b$  in Figure 1). In other words, inertia measures, in relative terms, how close post-completion holdings are to the fully passive benchmark, as opposed to the fully active benchmark. In the specific example shown in Figure 1, inertia at completion is 44.9 percent ( $44.9 = 32 - 14$  divided by  $53 - 14$ ).

## 2. Individual investor inertia

Table 2 presents our analysis of inertia among individual investors. This analysis is restricted to the 305 stock-swap mergers that have at least one target-only investor during the interval for which we have the brokerage-firm data. For each deal, we begin with all investors,

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<sup>19</sup> In the institutional data, we calculate the fully passive and fully active benchmarks by examining the behavior of investors in the same set of target stocks in the period beginning twelve quarters before merger announcement. For the purpose of computing the benchmark, we again focus on those target investors who have no initial position in the acquirer. The fully passive benchmark is these investors’ propensity to close out their position in the *target* over various horizons, and the fully active benchmark is their propensity to establish a new position in the *acquirer* over various horizons. We take a similar approach with individual investors, calculating the benchmarks in 1991 on all deals that are announced in 1993 or later.

who, at the month-end prior to the transaction, hold the target but not the acquirer. We then track the holdings of this set of individuals over the period from announcement to completion. Each of the first nine columns in the table isolates deals with a fixed length of time to completion (one month, two months, etc.) and the final column presents aggregated results for all 305 deals.

For example, the first column of Table 2 shows that we have 21 mergers with a time to completion of one month.<sup>20</sup> In these transactions, 11.4 percent of pre-merger individual investors close out their positions in the target *before* completion; thus pre-completion holdings are 88.6 percent ( $88.6 = 100 - 11.4$ ). Of the remaining 88.6 percent, another 26.3 percent close out their positions within three months *after* completion, leaving net post-completion holdings at 65.3 percent ( $65.3 = 88.6 \cdot (1 - 0.263)$ ) four months after we start tracking the holdings of the target-only investors in these 21 mergers.

This post-completion figure of 65.3 can then be compared to the fully passive four-month benchmark of 87.3 percent, which means that, in a typical non-merger-affected, four-month period for the full set of target stocks, 12.7 percent of individual investors close out their positions. It can also be compared to the fully active four-month benchmark of 0.3 percent, which means that in a typical non-merger-affected, four-month period for the full set of target stocks, only 0.3 percent of individual investors in the target who do not initially have a position in the acquirer open a new position in the acquirer.<sup>21</sup> Putting it all together, we calculate inertia of 74.7 percent for the subset of mergers that take one month from announcement to completion ( $74.7 = 65.3 - 0.3$  divided by  $87.3 - 0.3$ ).

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<sup>20</sup> Specifically, a merger is considered to have a time to completion of one month if it is announced at any time during one calendar month and completed at any time during the next.

<sup>21</sup> The finding that the lower benchmark is so low for individual investors is a consequence of the low diversification of individual investors. The median investor in our dataset holds two stocks at any given time.

The results are very similar in each of the subsequent columns, representing deals that take two or more months to move from announcement to completion. The final column aggregates across these different samples, weighting each one by the number of target-only investors. On average, for individuals, post-completion holdings are 64.1 percent, and inertia is 78.3 percent. Thus, loosely speaking, we estimate that about 80 percent of individuals effectively sleep through mergers in which they are shareholders in the target firm.

This is a rough number, and one can certainly argue with the details of our benchmarking methodology. However, it is important to note that our basic conclusion is likely to be robust: the vast majority of individual target-firm investors do not react to a merger by unloading their shares. Figure 2 highlights the conclusions in Table 2 graphically. Across all merger horizons, the ownership percentage falls at a similar rate to that of the passive benchmark. Essentially all horizons are within ten percent of this upper benchmark. Given that the active benchmark never gets above one percent, it is clear that the inertia among individual target investors is substantial.

### **3. Institutional investor inertia**

Table 3 undertakes an analogous exercise for institutional investors. Because of the broader coverage of the CDA Spectrum data, we are able to include 1,797 stock-swap mergers in this analysis. This is the subset of mergers that have at least one target-only investor between the second quarter of 1980 and the fourth quarter of 2002. Because of the SEC 13-F reporting requirements, we are forced to look at things on a quarterly, as opposed to monthly basis.

Aggregating across all transactions, we find that 30.0 percent of pre-merger institutional investors hold on to their positions through completion, substantially less than the corresponding figure of 64.1 percent for individuals. The baseline rate of turnover is also higher. However, the

first effect dominates and inertia is significantly lower, at 32.3 percent for institutions, as compared to 78.3 percent for individuals. In contrast to the individual investor turnover in Figure 2, institutional holdings in Figure 3 decline at significantly faster rate than the passive benchmark across all merger completion horizons. By the conclusion of most of the mergers, the institutional holdings lines are all closer to the active benchmark than to the passive one. Nonetheless, even the institutional holdings are considerably higher than the active benchmark, suggesting, under the assumptions of the model, that a stock swap is a more attractive way to raise equity than an SEO regardless of the target's ownership structure.

Undoubtedly though, institutions are much less passive than individuals. We exploit this fact in what follows, using the fraction of institutional investors in the target firm as a proxy for the parameter  $\alpha$  in our model. This allows us to undertake direct tests of Propositions 2 and 3.

#### **4. The determinants of inertia**

Before proceeding to Propositions 1 and 2, we consider whether there is an identifiable rational component to investor inertia. In particular, capital gains taxes might encourage investors to hold on to appreciated stock in a merger. Or, if the target is large relative to the acquirer, a target investor with no revealed preference for the acquirer may rationally want to hold on to the shares of the merged company, because its value is determined to a large extent by the prospects of the original target assets.

In Table 4, we present the pre- and post-completion turnover of target-only investors calculated for various subsamples of stock-swap mergers. In Panels A and D, we address the tax hypothesis by splitting the individual and institutional-investor samples according to the pre-announcement return of the target, calculated as the cumulative return over the two-year period



ending one month prior to announcement. Pre-announcement returns appear to be weakly related to the inertia of individual investors: target investors are somewhat less willing to dump shares that have appreciated in value by more than 20 percent. However, the differential across groups is quite modest.<sup>22</sup> Even mergers where the target has recently declined in value also involve considerable inertia; indeed, the inertia statistic of 76.5 percent for this group is close to the full-sample value of 78.3 percent. With institutions, there is no evidence at all to suggest that capital-gains taxes matter for our measure of inertia. Both raw post-completion holdings and our inertia statistic are at their highest among those firms with negative pre-announcement returns, the group where tax considerations would suggest that there should be the *least* inertia.

A somewhat sharper test of the tax hypothesis is contained in Panel C. Here we split the individual-investor sample according to whether or not the account in question is taxable or tax-exempt. As can be seen, the differences in inertia across the two subsamples are negligible, with inertia statistics of 78.4 percent and 81.4 percent for taxable and non-taxable accounts respectively. Thus the evidence strongly rejects the proposition that inertia is a rational response to capital-gains-tax considerations.

Panels B and E split the samples based on the relative size of the target, calculated as the target's market capitalization divided by the total market capitalization of the target and acquirer. Individuals exhibit uniformly high inertia in mergers of widely varying relative size. Institutions, on the other hand, appear slightly less passive when the target is large, accounting for at least 25 percent of combined firm value. Again, this result goes in the opposite direction relative to a simple rational story: mergers with large targets represent situations where the

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<sup>22</sup> Two effects combine to make the inertia statistic greater among targets with large positive returns. First, post-completion holdings are greater, consistent with a capital-gains tax story. And second, the upper (fully active) benchmark is lower, consistent with the disposition effect (Shefrin and Statman (1985)). Thus, if anything, the modest differences in inertia that we document likely overstate the pure impact of capital-gains-tax considerations.

combined firm's assets ought to appear most attractive to those investors who initially found the target worth owning, so such deals should be expected to generate more, not less inertia.

### *C. Acquirer Returns and Volume Around Merger Announcements*

Having established the main premise of the model – that investors exhibit inertia – we now consider its empirical implications. In particular, Tables 5 through 7 present our tests of Propositions 1 and 2. The first part of Proposition 1 is that target-investor inertia leads to less negative acquirer announcement returns. A corollary laid out in Proposition 2 is that this effect works through volume, so that target-investor inertia leads to lower acquirer announcement volume at the same time. The third part of Proposition 1 is that the impact of target-investor inertia interacts with the slope of the demand curve for acquirer stock. The steeper is this demand curve, the greater should be the impact of inertia on acquirer announcement returns.

#### **1. Institutional ownership and acquirer announcement returns**

In Panel A of Table 5, we focus on stock-for-stock mergers, and regress acquirer announcement returns on both target institutional ownership, and non-overlapping target institutional ownership, as defined above. The first regression uses the raw measure of institutional ownership as our proxy for investor wakefulness  $\alpha$ , and has no other controls, other than year fixed effects. This univariate specification generates a coefficient of  $-3.86$  on institutional ownership, which is strongly statistically significant (t-statistic = 4.72), consistent with Proposition 1. In economic terms, this coefficient implies that a two-standard deviation increase in target institutional ownership reduces the acquirer's announcement return by 1.76

percentage points, taking it from its unconditional mean value of  $-2.22$  percent down to  $-3.98$  percent.

The next regression again uses the raw measure of institutional ownership, but adds a variety of controls described in Table 1: acquirer and target market capitalization (both in logs); acquirer and target leverage; acquirer and target market-to-book; the target announcement return; and an indicator variable equal to one if the two firms are in the same industry. Many of these (the size, leverage, market-to-book, and relatedness variables) are commonly used in regressions to explain acquirer announcement returns.<sup>23</sup> We add the target announcement return to the list because one potential competing explanation for the effect of institutional ownership has to do with bargaining power. Perhaps institutional blockholders in the target are able to extract a better price from the acquiring firm, leading to lower acquirer announcement returns; if so, the target announcement return should control for this effect. However, as it turns out, this variable has a significant, but *positive*, relationship with acquirer returns.<sup>24</sup>

In any case, adding the full battery of controls has only a modest impact on the coefficient on institutional ownership: it goes from  $-3.86$  to  $-2.62$ , and remains statistically significant, with a t-statistic of 2.19. In this specification, a two-standard-deviation increase in target institutional ownership reduces the acquirer return by 1.19 percentage points.

The next two columns of Panel A are analogous to the previous two, except that we replace target institutional ownership with non-overlapping target institutional ownership. This latter variable is arguably a more precise measure of the wakefulness parameter  $\alpha$  in our model,

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<sup>23</sup> See, e.g., Lang, Stulz, and Walkling (1989) and Morck, Shleifer, and Vishny (1990b). Other typical merger-related controls are not included because of the nature of our sample. The acquirer attitude (see Schwert (2000) for a discussion of hostility) is always friendly in our sample of stock swaps, and by definition the form of payment (see Andrade, Mitchell, and Stafford (2001)) is stock and the accounting treatment is pooling (see Martinez-Jerez (2004)).

<sup>24</sup> This is unlikely to be a causal relationship. More plausibly, some mergers are just better than others in which case both acquirer and target returns are high.

because we expect an alert target shareholder to be most likely to sell shares in the acquirer if he did not own any such shares prior to the merger announcement, and hence has not demonstrated a high valuation for acquirer assets. Interestingly, this redefinition of target institutional ownership leads to coefficient estimates that are markedly higher in absolute value: they are now  $-8.14$  and  $-5.47$  in the no-controls and full-controls specifications, respectively. We view this pattern as particularly supportive of our model, because it is hard to think of alternative hypotheses which would suggest a similar outcome.

Panel B of Table 5 is an exact replica of Panel A, except that the sample includes cash mergers instead of stock mergers. This is effectively a placebo check: according to our theory, target institutional ownership should be irrelevant in cash deals. In contrast, some competing explanations for the results in Panel A suggest a similar pattern across stock and cash mergers. For example, if high institutional ownership of the target leads to low bidder returns through some sort of enhanced-bargaining effect, this should work similarly for both stock and cash deals. However, as can be seen in Panel B, there is no discernible impact of target institutional ownership in the cash-merger sample: the coefficient of interest is never close to statistically significant, and indeed is actually positive in three out of four cases. Thus cash mergers seem to be fundamentally different from stock mergers on this dimension, consistent with our model.

## **2. Institutional ownership and acquirer announcement volume**

The model makes the ancillary prediction that in stock deals, the impact of shareholder inertia on acquirer announcement returns works through trading volume, as in Proposition 2. In particular, target shareholders who are awake sell out on announcement. Merger arbitrageurs buy these shares and short the acquirer, closing their positions when the merger is successfully

completed.<sup>25</sup> Examining volume also serves as yet another check on alternative hypotheses linking target institutional ownership and acquirer returns. Again, if target institutional ownership affects acquirer returns through a mechanism like bargaining power, we would not expect it to also influence volume simultaneously.

Table 6 repeats the analysis in Table 5, replacing acquirer returns as the dependent variable with the average daily turnover in the acquirer over the five-day window surrounding the merger announcement. We keep all the same right-hand-side variables as before, and also add “normal” trading volume – defined as the average daily turnover in the acquirer over the 60-day period starting 90 days before the announcement – as an additional control.

Across all four specifications in the stock-merger sample in Panel A, the results are uniformly supportive of Proposition 2. The coefficients on raw target institutional ownership are 0.58 and 1.07 in the specifications with and without controls, respectively, and are strongly statistically significant in both cases. When we use non-overlapping target institutional ownership instead, the coefficients again rise in absolute value, to 0.86 and 1.71 respectively. And as before, the implied economic effects are substantial: a two-standard deviation increase in target institutional ownership increases the average acquirer’s daily turnover during the announcement period from a mean of 1.00 percent to between 1.26 and 1.49 percent, or by between 25 and 50 percent. In Panel B, we see that for cash mergers, target institutional ownership has little effect on turnover, just as it has no effect on returns: the coefficients are in all cases much smaller than in Panel A, and, with one exception, not statistically significant. Again, this is just what one would expect based on our model.

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<sup>25</sup> Empirically, this is similar in spirit to Mitchell, Pulvino, and Stafford (2004), who look at situations where merger arbitrageurs are likely to initiate short positions in the acquiring firm. Here, target institutional ownership is the proxy for the selling pressure that pushes down the price of the acquirer.

### 3. Demand curve interactions

There are two ingredients to our model: inertia among target shareholders, and a downward-sloping demand curve for acquirer shares. Thus, as formulated in the third part of Proposition 1, we expect that our results for target institutional ownership will be strongest among acquirer firms with steeply-sloped demand curves. To operationalize this hypothesis, we employ two different proxies for the slope of the demand curve. The first aims to measure the difference of opinion among investors with respect to acquirer value, in the spirit of the parameter  $H$  in the model. Specifically, we follow Moeller, Schlingemann and Stulz (2004), and calculate the standard deviation of all outstanding analysts' forecasts for long-run growth.

The second proxy is the non-industry idiosyncratic risk of the acquirer. We compute this as the standard deviation of the residuals from a regression of acquirer excess returns on the Fama-French factors ( $R_M$ ,  $SMB$ ,  $HML$ ), and the matched 48-industry portfolio return. The premise here is as follows. In the presence of both differences of opinion and risk aversion, an increase in idiosyncratic risk makes the demand curve steeper, because it reduces the size of the position that any one investor with a given valuation is willing to take on. Although this effect is absent from our model (which, for simplicity, uses wealth constraints instead of risk aversion to generate the shape of the demand curve) it is formalized in, e.g., Chen, Hong, and Stein (2002).<sup>26</sup> Moreover, Wurgler and Zhuravskaya (2002) provide empirical validation for the idea of using idiosyncratic risk as a proxy for demand-curve slope, showing that the impact of S&P 500 index inclusion on stock prices is increasing in the idiosyncratic risk of the included firm.

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<sup>26</sup> In particular, see equation 4 of Chen, Hong, and Stein (2002). This equation demonstrates that the sensitivity of price to supply shocks is larger in absolute value when the risk tolerance of the heterogeneous buyers is small. An increase in idiosyncratic risk has an analogous effect. See also Novy-Marx (2004) for more on the general argument that risk aversion makes demand curves slope downward, and hence creates the appearance of illiquidity.

Table 7 presents regressions of acquirer announcement returns (in stock deals only) on target institutional ownership, our measures of the slope of the demand curve facing the acquirer, and the product of the two variables. In Panel A, we use the raw measure of institutional ownership; in Panel B we use non-overlapping institutional ownership. In either case, our theory suggests that the interaction term should attract a negative coefficient. The first two regressions in each panel employ dispersion in analyst forecasts as the proxy for the steepness of the acquirer's demand curve, while the second two use idiosyncratic risk.

The results in Table 7 provide further corroboration of the model. The interaction of the acquirer demand-curve proxies and target institutional ownership is negative in all eight specifications shown in the table, and significant (at the ten percent level or better) in the four specifications in Panel A that use raw institutional ownership. In Panel B, with non-overlapping institutional ownership, the interaction coefficients are generally as big or bigger, but the standard errors are larger, perhaps because non-overlapping ownership has less cross-sectional variation.

Moreover, each of the eight specifications implies economically meaningful interaction effects. The demand-curve proxies are standardized to have zero mean and unit variance. So the fact that the coefficients on target institutional ownership and on the interaction term are of roughly the same magnitude in all specifications means that a one-standard-deviation increase in the steepness of the demand curve roughly doubles the effect of institutional ownership on acquirer returns.

To sum up, higher target institutional ownership leads to more negative announcement effects for the acquirer. Our additional tests help pin down the mechanism. Institutional holders of the target exhibit less inertia and so are more likely to sell their shares on the announcement of

a merger. This leads to more selling pressure in the acquirer and a negative announcement effect that is strongest for those acquiring firms with steeply-sloped demand curves.

#### **IV. Further Implications**

In addition to its implications for merger announcement effects, our model makes two more speculative predictions for corporate finance. All else equal, inertia makes equity more attractive than cash as consideration in a merger or acquisition, and it makes equity issued in the context of a merger more attractive than equity issued in an SEO.<sup>27</sup>

##### *A. Financing Choice*

The first hypothesis is that stock-for-stock mergers are more attractive to an acquiring firm when a greater fraction of target shareholders are asleep. This is a natural prediction in a more general version of the model, where the acquiring firm can use cash or shares as consideration. To make the financing choice interesting, we need to introduce a cost of using cash, like the expected costs of financial distress. With this extra ingredient, it is straightforward to show that an increase in target shareholder inertia makes stock financing relatively more attractive.

Using the entire sample of cash and stock mergers, we run a probit regression in which the dependent variable is a dummy that equals one if the consideration in the deal is stock (and zero if it is cash), and in which the independent variables include either target institutional

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<sup>27</sup> Note that in formulating these hypotheses, we are taking the identity of the target firm as exogenous—i.e., as determined by factors outside of our model, such as the quality of the match between the acquirer and the target. It is tempting to add the additional prediction that a firm with a greater fraction of sleeping shareholders is more likely to become a target in the first place. However, the same inertia that improves the terms for the acquiring firm may also lead to target management entrenchment.



ownership, or non-overlapping target institutional ownership, as well as the controls used in previous tables. As can be seen in Table 8, either measure of institutional ownership has a strong negative influence on the probability of a merger being done with stock. We report the coefficients as percentage effects evaluated at the mean of each independent variable. So, a two-standard deviation increase in target institutional ownership reduces the probability of a stock offering by 9.7 percent from a mean of 61.9 percent, and a two-standard deviation increase in non-overlapping institutional ownership reduces the probability by 17.0 percent. Both are statistically significant as is the difference between the two estimates. Given that institutional investors are less likely to be sleepers than individuals, this pattern fits with our hypothesis.

Of course, there are alternative interpretations for the results in Table 8. Maybe some institutions simply have a relative preference – as compared to individual investors – for receiving cash as opposed to stock in a merger transaction, and acquirers cater to this preference in order to lower the required merger premium. We cannot directly refute this alternative story, though it would seem to have a hard time rationalizing the fact that our results are actually considerably stronger for the non-overlapping measure of institutional ownership that is most directly suggested by our theoretical model. Moreover, because acquirer shares can always be converted to cash in the market, an institutional preference for cash requires that institutions expect significant transaction costs in exiting positions in the acquirer. The focus of our model is precisely on the price pressure component of these transaction costs.

### *B. Equity Financing in Mergers and SEOs*

Our theory also has something to say about why equity financing might be more common in the context of mergers than it is with greenfield investment (Fama and French (2004)). As

outlined in the introduction, this point is easiest to see by thinking of a firm with an exogenously specified growth strategy that, over the next year, will entail one acquisition, and one major new greenfield investment, with the two transactions being of roughly similar size. Assume that, in order to keep its capital structure in balance, the firm will need to finance one of these two investments with an equity issue – i.e., either the merger will have to be stock-for-stock, or the greenfield investment will have to be accompanied by an SEO.<sup>28</sup>

We have seen that an SEO effectively amounts to a limiting case of our model with no inertia, since all shares must be sold off in the market, and none can simply be placed in the hands of inertial investors. Thus the SEO will be associated with a more negative price impact than the stock-for-stock merger, all else equal, and hence will tend to be less attractive. So we would expect the firm to finance the merger with a stock swap, but to finance the greenfield investment with cash. As noted above, this is a prediction that fits closely with the empirical results of Fama and French (2004).

Pushing the logic further, our model also suggests that the relative preference for doing the merger as a stock swap as opposed to financing the greenfield investment with an SEO will be stronger when: (i) the difference of opinion  $H$  among acquiring-firm investors is greater; (ii) more target shareholders are asleep; and (iii) the scale of the two investments in question is larger.

Although we do not test any of these hypotheses formally, we can offer a bit of suggestive evidence. We have done some further comparisons of merger and SEO activity that

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<sup>28</sup> A firm that considers its stock to be overvalued, and that wants to exploit this overvaluation, can be thought of as “needing” to finance some of its investments with equity issues in this sense. Our model then explains why stock-for-stock mergers are a more attractive option for such an overvalued firm than SEO-financed capital expenditures, even when the assets ultimately acquired would be similar in either case. This contrasts with Shleifer and Vishny (2003) who argue that stock-for-stock mergers are motivated by overvaluation, but do not explain why they are any better an outlet for overvalued acquirers than SEO-financed capital expenditures.

speak to scale effects. Based on a sample of 6,526 SEOs and 2,040 stock-for-stock mergers of public companies over the period 1990-2003, we find that not only do mergers raise substantially more total equity financing than SEOs (\$2,559 billion vs. \$968 billion), but that the largest transactions – both in absolute terms and relative to issuer size – are much more likely to be mergers than SEOs. For example, about 53 percent of stock-for-stock mergers are for more than \$100 million, while only 33 percent of SEOs are. Alternatively, 23 percent of mergers are for more than 50 percent of the issuer’s market value, while only 10 percent of SEOs are.<sup>29</sup>

Again, we stress that these stylized facts do not represent a decisive test – they are also consistent with other interpretations, such as mergers having an advantage over greenfield investment when it comes to big projects where time-to-build considerations are likely to be important. Nevertheless, they do fit nicely with the model’s implication that the relative appeal of a stock-for-stock merger is greatest when the firm’s growth plan is such that it needs to raise a very large amount of equity financing.

## V. Conclusions

Most people are reluctant to make active decisions. Instead, they tend to follow the path of least resistance, accepting defaults. This has important implications for corporate finance. Raising equity in an SEO requires investors to *actively buy* the shares of the issuing firm. The default option is to buy no shares. By contrast, raising equity in a stock-swap merger works well to the extent that target investors *do not actively opt out of holding* the shares of the acquiring firm. The default option is to accept the shares of the acquiring firm as consideration for shares

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<sup>29</sup> The numbers in this paragraph are based on an analysis of transactions in the Thomson Financial mergers and acquisitions and seasoned equity offerings databases.

in target. We find that this sort of inertia is a pervasive aspect of investor behavior. Individuals accept the default roughly 80 percent of the time, and institutions accept it a third of the time.

In a classical stock market with horizontal demand curves, inertial behavior of this kind would be irrelevant for prices. However, combining inertia with a downward-sloping demand curve makes the price impact of a stock-swap merger relatively small, and the terms for existing shareholders better as a result. We test this idea within the sample of stock-swap mergers, using institutional ownership to proxy for low investor inertia. Consistent with the theory, acquirer announcement returns are more negative when inertia is low. The broader conclusion for corporate finance is that, when firms face downward-sloping demand curves – as numerous studies suggest – stock-swap mergers may play a particularly important role in supporting a strategy of rapid, equity-financed growth.

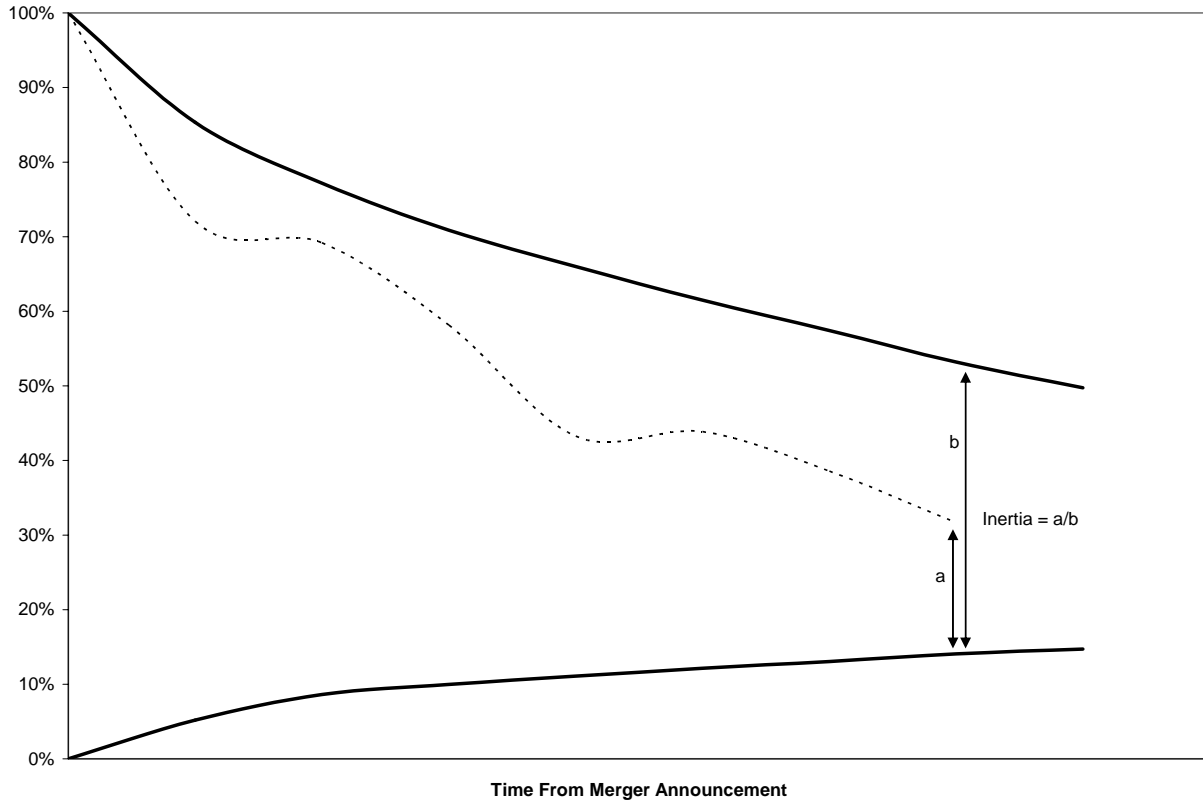
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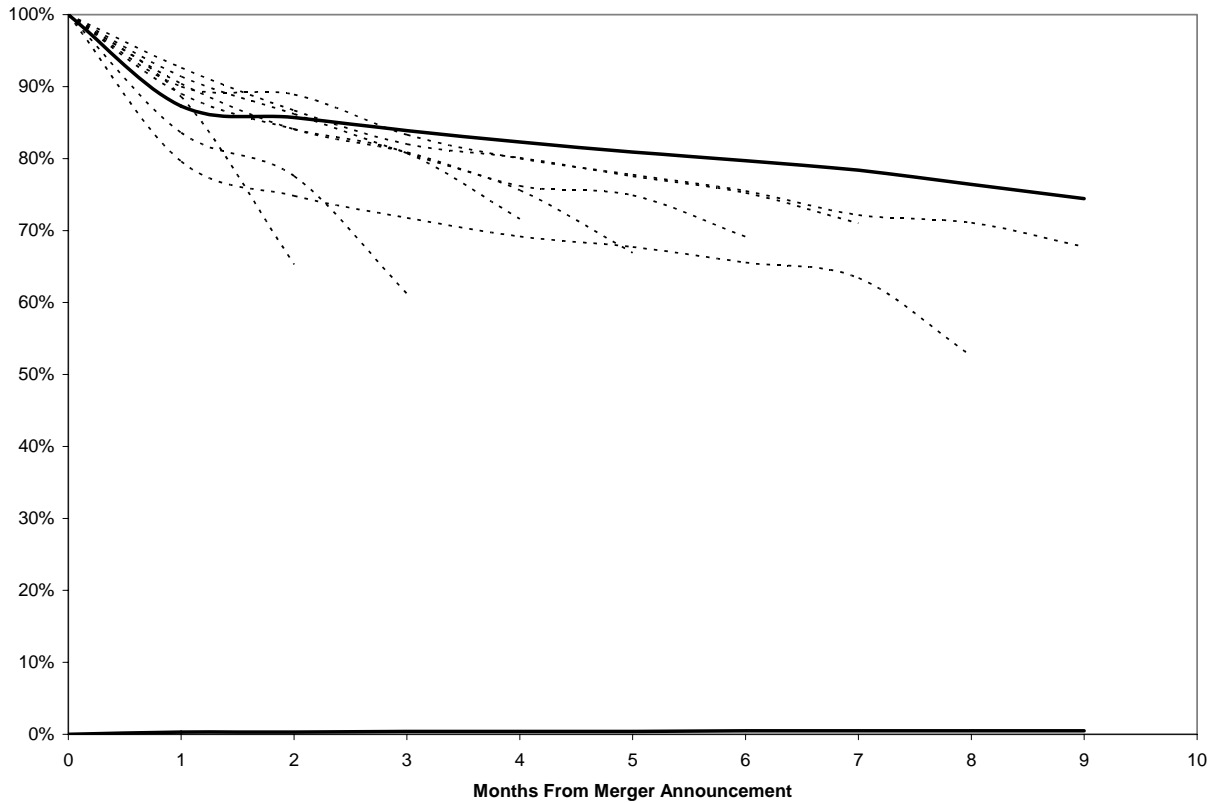
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**Figure 1. Calculating inertia.** To calculate the fraction of investors that are passive with respect to a merger transaction, we compare their holdings to benchmark holdings levels (in bold). The upper benchmark reflects the fraction of investors that, during a non-merger period, held a given stock in month zero and continued to hold it in subsequent months. The lower benchmark reflects the fraction of investors that, during a non-merger period, did not hold a given stock in month zero but did hold it during subsequent months. These benchmarks can be compared to the fraction of investors that, at the time of a merger announcement, held the target but not the acquirer, and who continue to hold the target prior to completion or who hold the acquirer after the merger is completed.

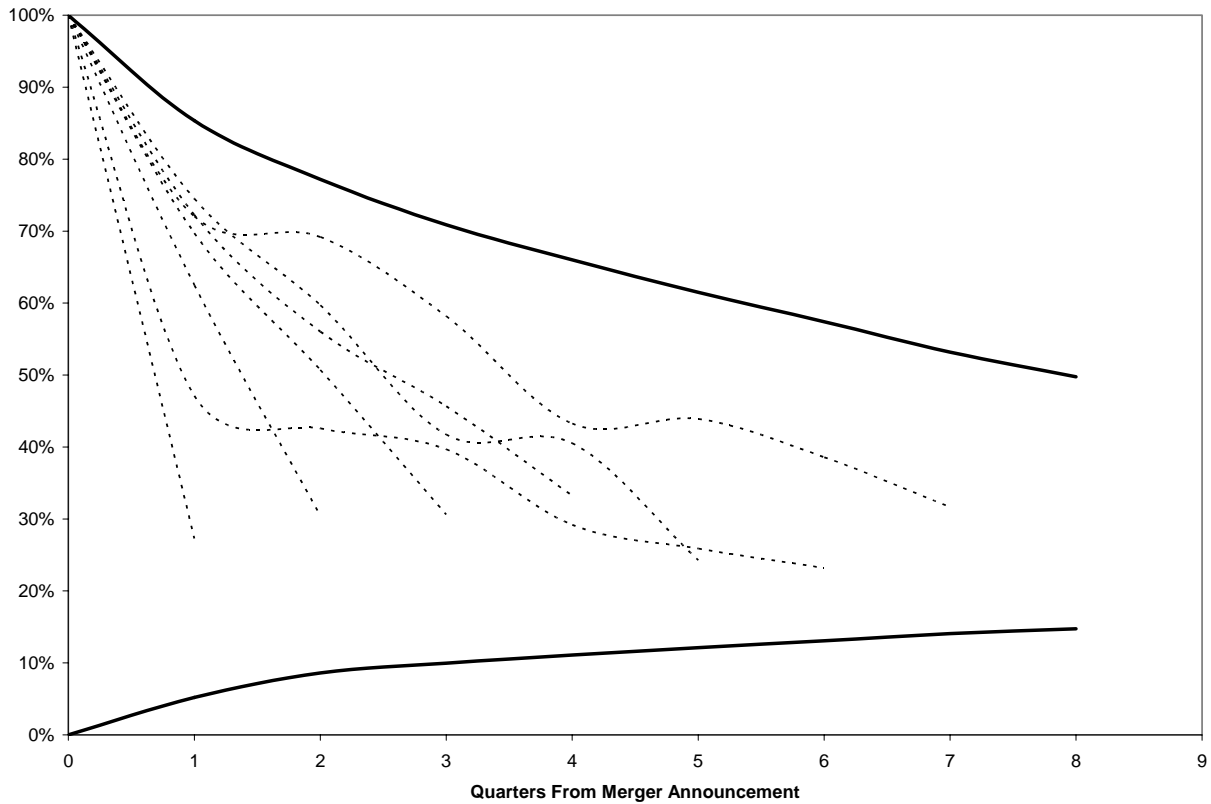




**Figure 2. Turnover around merger announcements: Individual holdings.** We identify situations in a database of individual investor holdings at a large discount broker from 1991 to 1996 where the investor has a position in a target but not its acquirer in the month ending prior to the announcement of a stock-swap merger in the *Wall Street Journal*. The figure includes 305 successful stock-swap mergers that are completed within nine months involving a public acquirer and with at least one matched individual investor with a position in the target but not the acquirer. Each dashed line tracks the percentage of the original individual investors that still has a position in the target at each month end prior to completion and that has a position in the acquirer two full months following the completion of the stock swap. We split the sample, plotting one line for each number of months elapsed between the announcement and completion of a merger. The solid lines benchmark the individual investor holdings. The upper benchmark holdings reflects the corresponding turnover for situations in the individual investor database where an individual owns shares in a target but not its acquirer in 1991 for all deals that are announced in 1993 or later. The lower benchmark starts with the same situations and tracks the purchase of the acquirer shares.



**Figure 3. Turnover around merger announcements: Institutional holdings.** We identify situations in the CDA Spectrum Institutional Holdings database where an institution has a position in a target but not its acquirer in the quarter ending prior to the announcement of a stock-swap merger in the *Wall Street Journal*. The figure includes 1,789 successful stock-swap mergers that are completed within six quarters, involving a public acquirer and with at least one matched institution with a position in the target but not the acquirer. Each dashed line tracks the percentage of the original institutions that still has a position in the target at each quarter end prior to completion and that has a position in the acquirer at the quarter end following the completion of the stock swap. We split the sample, plotting one line for each number of quarters elapsed between the announcement and completion of a merger. The solid lines benchmark the institutional holdings. The upper benchmark holdings reflects the corresponding turnover for situations in the CDA Spectrum Institutional Holdings database where an institution owns shares in a target but not its acquirer in the quarter ending *twelve quarters prior* to announcement. The lower benchmark starts with the same situations twelve quarters prior to announcement and tracks the purchase of the acquirer shares.



**Table 1. Merger summary statistics.** The sample includes successful stock-swap and cash mergers with CRSP targets announced between 1980 Q2 and 2000 Q4, involving a CRSP acquirer with a matched announcement return. Institutional ownership is summarized in Panel A. Target institutional ownership (IO) is totaled from the CDA Spectrum Institutional Holdings database for the quarter prior to the announcement of the merger and expressed as a percentage of shares outstanding. Non-overlapping target IO includes only those institutions that own no shares of the acquiring firm. This is expressed as a percentage of both shares outstanding and total target IO. The deal characteristics in Panel B are from CRSP and Compustat. Cash deal is equal to one when the consideration is cash and zero when the consideration is stock. Acquirer and target size are equal to price times shares outstanding from CRSP. Relative size is equal to target size expressed as a percentage of total target and acquirer size. Acquirer and target leverage are equal to interest bearing debt (9+34) from Compustat expressed as a percentage of book assets. Acquirer and target market to book are equal to book assets (6) minus book equity (216-130+35) from Compustat plus price times shares outstanding from CRSP all divided by book assets. Same industry is an indicator variable equal to one if the target and the acquirer are included in the same Fama and French 48-industry grouping. The stock market data in Panel C are from CRSP. The acquirer and target announcement returns are the return in excess of the value-weighted market over a five-day window surrounding the announcement of the merger. The acquirer announcement volume is the average daily volume over a five-day window surrounding the announcement of the merger expressed as a percentage of shares outstanding. Normal volume is the average daily volume over a 60-day window starting 90 trading days before the announcement of the merger and expressed as a percentage of shares outstanding. The acquirer demand curve proxies in Panel D are from CRSP and IBES. The dispersion in analyst forecasts is the standard deviation of all outstanding long-term growth forecasts. Idiosyncratic risk is the standard deviation of the residuals from a regression of acquirer excess returns on the Fama-French Benchmark Factors (Rm, SMB, HML), and the matched 48 industry portfolio return. All factor and portfolio returns were obtained from Ken French's website. All variables are Winsorized at the first and 99<sup>th</sup> percentiles.

	Summary Statistics, 1980Q2-2000Q4					
	N	Mean	Median	SD	Min	Max
<b>Panel A. Institutional Ownership</b>						
Target IO (% Total)	3,003	26.36	20.47	22.72	0.00	85.19
Non-Overlapping Target IO (% Total)	3,003	13.10	8.92	13.62	0.00	61.30
Non-Overlapping Target IO (% IO)	2,829	51.01	50.56	30.86	0.00	100.00
<b>Panel B. Deal Characteristics</b>						
Cash Deal	3,003	0.38	0.00	0.49	0.00	1.00
Acquirer Size (\$M)	3,003	7,224	1,102	21,000	5	146,000
Target Size (\$M)	3,003	491	106	1,195	0	7,680
Relative Size (%)	3,000	17.67	11.25	18.27	0.00	85.77
Acquirer Leverage (%)	2,820	21.92	19.30	16.64	0.00	78.47
Target Leverage (%)	2,516	21.73	17.35	20.34	0.00	84.93
Acquirer M/B	2,839	2.42	1.30	3.26	0.63	22.04
Target M/B	2,545	2.01	1.30	2.05	0.45	14.01
Same Industry	3,003	0.57	1.00	0.50	0.00	1.00
<b>Panel C. Stock Market Data</b>						
Acquirer Announcement Return (%)	3,003	-1.16	-0.94	7.29	-24.09	20.82
Target Announcement Return (%)	3,003	18.65	14.71	21.78	-24.50	97.37
Acquirer Announcement Volume (%)	2,971	0.81	0.38	1.14	0.01	6.65
Normal Volume (%)	2,937	0.49	0.29	0.60	0.01	3.53
<b>Panel D. Acquirer Demand Curve</b>						
Dispersion in Analyst Forecasts (%)	1,606	3.27	2.50	3.07	0.00	60.90
Idiosyncratic Risk (%)	2,855	8.06	6.71	4.66	2.54	27.59

**Table 2. Turnover around merger announcements: Individual holdings.** We identify situations in a database of individual investor holdings at a large discount broker from 1991 to 1996 where the investor has a position in a target but not its acquirer in the month ending prior to the announcement of a stock-swap merger in the *Wall Street Journal*. The first row shows the number of successful stock-swap mergers, involving a public acquirer and with at least one matched individual investor with a position in the target but not the acquirer. We split the sample according to the number of months elapsed between the announcement and completion of a merger. The following ten rows show percentage turnover in this set of target positions over the nine months following the merger announcement and in the three-month period that the merger is successfully completed. Turnover in the quarter that the merger is completed occurs when one of the original investors still has a position in the target in the month prior to completion but does not have a position in the acquirer three months later. The next two rows compound the pre-completion and completion turnover to compute the percentage of the original investors that still has a position in the target at the month end prior to completion and that has a position in the acquirer three months later. The last two rows benchmark the post-completion holdings. The upper benchmark holdings reflects the corresponding turnover for situations in the individual investor database where an individual owns shares in a target but not its acquirer in 1991 for all deals that are announced in 1993 or later. The lower benchmark starts with the same situations and tracks the purchase of the acquirer shares. The final row reports the difference between post-completion holdings and the lower benchmark expressed as a percentage of the difference between the upper and lower benchmarks.

	Total Months to Completion									
	1	2	3	4	5	6	7	8	9+	All
Number of Successful Stock Swaps	21	76	56	44	39	32	12	8	17	305
Announcement+1 Turnover (%)	11.4	16.4	11.0	8.6	9.6	7.4	20.4	10.0	6.5	12.1
Announcement+2 Turnover (%)		7.2	5.5	5.7	7.0	6.4	6.0	1.2	5.0	6.1
Announcement+3 Turnover (%)			4.0	6.3	3.8	5.4	4.1	6.3	1.4	4.4
Announcement+4 Turnover (%)				6.4	5.8	2.3	3.6	4.0	5.2	5.0
Announcement+5 Turnover (%)					1.7	3.2	2.1	2.8	4.4	2.9
Announcement+6 Turnover (%)						3.0	3.2	2.9	7.7	4.8
Announcement+7 Turnover (%)							3.3	4.4	10.1	6.5
Announcement+8 Turnover (%)								1.5	3.7	3.4
Announcement+9 Turnover (%)									3.1	3.1
Completion Turnover (%)	26.3	21.1	11.3	11.5	7.7	5.6	17.1	4.7	6.9	14.0
Pre-Completion Holdings (%)	88.6	77.6	80.8	75.6	74.9	75.2	63.4	71.1	56.4	74.7
Post-Completion Holdings (%)	65.3	61.2	71.6	66.9	69.1	70.9	52.6	67.8	57.2	64.1
Upper Benchmark Holdings (%)	87.3	85.7	83.9	82.3	80.9	79.7	78.4	76.4	70.7	81.7
Lower Benchmark Holdings (%)	0.3	0.3	0.4	0.4	0.4	0.5	0.5	0.5	0.6	0.6
Inertia (%)	74.7	71.3	85.3	81.2	85.3	88.9	66.9	88.7	80.7	78.3

**Table 3. Turnover around merger announcements: Institutional holdings.** We identify situations in the CDA Spectrum Institutional Holdings database where an institution has a position in a target but not its acquirer in the quarter ending prior to the announcement of a stock-swap merger in the *Wall Street Journal*. The first row shows the number of successful stock-swap mergers, involving a public acquirer and with at least one matched institution with a position in the target but not the acquirer. We split the sample according to the number of quarters elapsed between the announcement and completion of a merger. The following five rows show percentage turnover in this set of target positions over the four quarters following the merger announcement and in the quarter that the merger is successfully completed. Turnover in the quarter that the merger is completed occurs when one of the original institutions still has a position in the target in the quarter prior to completion but does not have a position in the acquirer in the quarter after completion. The next two rows compound the pre-completion and completion turnover to compute the percentage of the original institutions that still has a position in the target at the quarter end prior to completion and that has a position in the acquirer at the quarter end following the completion of the stock swap. The last two rows benchmark the post-completion holdings. The upper benchmark holdings reflects the corresponding turnover for situations in the CDA Spectrum Institutional Holdings database where an institution owns shares in a target but not its acquirer in the quarter ending *twelve quarters prior* to announcement. The lower benchmark starts with the same situations twelve quarters prior to announcement and tracks the purchase of the acquirer shares. The final row reports the difference between post-completion holdings and the lower benchmark expressed as a percentage of the difference between the upper and lower benchmarks.

	Total Quarters to Completion					All
	0	1	2	3	4+	
Number of Successful Stock Swaps	85	862	609	166	75	1,797
Announcement+1 Turnover (%)		37.5	30.3	27.8	36.0	34.4
Announcement+2 Turnover (%)			27.3	22.5	8.3	23.8
Announcement+3 Turnover (%)				18.5	12.9	17.5
Announcement+4 Turnover (%)					11.0	16.7
Completion Turnover (%)	72.7	51.3	39.6	27.2	26.6	43.6
Pre-Completion Holdings (%)		62.5	50.7	45.6	33.8	55.0
Post-Completion Holdings (%)	27.3	30.4	30.6	33.2	23.4	30.0
Upper Benchmark Holdings (%)	85.3	77.2	70.9	66.0	56.9	73.1
Lower Benchmark Holdings (%)	5.2	8.6	10.0	11.1	13.3	9.5
Inertia (%)	27.6	31.8	33.9	40.3	23.2	32.3

**Table 4. Turnover around merger announcements: Subsamples based on deal characteristics.** We repeat the analysis in Tables 2 and 3 for subsamples of stock swaps. See Tables 2 and 3 for details. In Panels A and D, we split the successful stock-swap mergers with at least one matched target position, into three groups according the return in the two years ending one month prior to the announcement of the merger in the *Wall Street Journal*. In Panels B and E, we split the sample according to the relative size of the target and acquiring firms. Relative size is equal to the target market capitalization (price times shares outstanding from CRSP) expressed as a percentage of the total market capitalization of the target and the acquirer. In Panel C, we split the sample of individuals according to whether or not any of their accounts are taxable.

	Pre-Completion	Post-Completion			
	Holdings	Holdings	Upper	Lower	Inertia
<b>Panel A. Individuals, Split on Target 2-Year Pre-Announcement Return</b>					
<0%	75.2	61.2	79.9	0.2	76.5
0-20%	83.3	71.3	78.9	0.3	90.3
>20%	74.9	65.6	70.1	0.4	93.5
<b>Panel B. Individuals, Split on Relative Size</b>					
<5%	77.1	62.7	74.8	0.3	83.8
5-25%	73.7	63.7	77.9	0.3	81.7
>25%	74.1	64.5	76.8	0.2	83.9
<b>Panel C. Individuals, Split on Tax Status</b>					
Taxable	67.9	60.0	76.7	0.2	78.4
Tax-exempt	78.1	63.3	78.2	0.2	81.2
<b>Panel D. Institutions, Split on Target 2-Year Pre-Announcement Return</b>					
<0%	53.5	33.1	73.4	9.4	37.0
0-20%	45.3	24.2	69.0	15.3	16.5
>20%	57.0	30.6	74.0	9.2	33.0
<b>Panel E. Institutions, Split on Relative Size</b>					
<5%	54.6	31.6	74.5	9.2	34.4
5-25%	50.8	33.8	75.3	13.1	33.3
>25%	58.2	27.8	73.4	9.4	28.8

**Table 5. Acquirer announcement returns.** Regressions of acquirer merger announcement returns on target institutional ownership and deal characteristics. Panel A shows results for stock swap mergers, and Panel B shows results for cash mergers. The acquirer announcement return is the return in excess of the value-weighted market over a five-day window surrounding the announcement of the merger. Target institutional ownership (IO) is totaled from the CDA Spectrum Institutional Holdings database for the quarter prior to the announcement of the merger and expressed as a percentage of shares outstanding. Non-overlapping target IO includes only those institutions that own no shares of the acquiring firm and is also expressed as a percentage of shares outstanding. The other deal characteristics are described in Table 1. All variables are Winsorized at the first and 99<sup>th</sup> percentiles. Announcement year fixed effects are included in all four specifications. Heteroskedasticity-robust t-statistics are reported in braces.

	Announcement Return (%)							
	Coef	[t]	Coef	[t]	Coef	[t]	Coef	[t]
<b>Panel A. Stock Deals</b>								
Target IO	-3.86	[-4.72]	-2.62	[-2.19]				
Non-Overlapping Target IO					-8.14	[-5.44]	-5.47	[-3.07]
log(Acquirer Size)			0.42	[2.86]			0.33	[2.21]
log(Target Size)			-0.58	[-2.78]			-0.62	[-3.46]
Acquirer Leverage			0.83	[0.51]			0.93	[0.58]
Target Leverage			0.92	[0.82]			0.90	[0.81]
Acquirer M/B			-0.27	[-3.04]			-0.27	[-3.03]
Target M/B			0.24	[1.51]			0.24	[1.48]
Target Announcement Return			0.06	[5.17]			0.06	[5.17]
Same Industry			-0.20	[-0.41]			-0.24	[-0.49]
Year FE		Yes		Yes		Yes		Yes
N		1,858		1,478		1,858		1,478
R <sup>2</sup>		0.03		0.08		0.03		0.08
<b>Panel B. Cash Deals</b>								
Target IO	-0.44	[-0.50]	1.50	[1.26]				
Non-Overlapping Target IO					0.91	[0.68]	1.30	[0.75]
Additional Controls		No		Yes		No		Yes
Year FE		Yes		Yes		Yes		Yes
N		1,145		903		1,145		903
R <sup>2</sup>		0.03		0.05		0.03		0.05

**Table 6. Acquirer announcement volume.** Regressions of acquirer merger announcement volume on normal volume, target institutional ownership, and deal characteristics. Panel A shows results for stock swap mergers, and Panel B shows results for cash mergers. The acquirer announcement volume is the average daily volume over a five-day window surrounding the announcement of the merger expressed as a percentage of shares outstanding. Normal volume is the average daily volume over a 60-day window starting 90 trading days before the announcement of the merger and expressed as a percentage of shares outstanding. Target institutional ownership (IO) is totaled from the CDA Spectrum Institutional Holdings database for the quarter prior to the announcement of the merger and expressed as a percentage of shares outstanding. Non-overlapping target IO includes only those institutions that own no shares of the acquiring firm and is also expressed as a percentage of shares outstanding. The other deal characteristics are described in Table 1. All variables are Winsorized at the first and 99<sup>th</sup> percentiles. Announcement year fixed effects are included in all four specifications. Heteroskedasticity-robust t-statistics are reported in braces.

	Acquirer Volume (%)							
	Coef	[t]	Coef	[t]	Coef	[t]	Coef	[t]
Panel A. Stock Deals								
Target IO	1.07	[10.94]	0.58	[4.51]				
Non-Overlapping Target IO					1.71	[7.91]	0.86	[3.67]
Normal Volume	1.35	[24.05]	1.27	[18.73]	1.35	[23.46]	1.27	[18.64]
log(Acquirer Size)			-0.19	[-11.68]			-0.18	[-10.95]
log(Target Size)			0.21	[10.25]			0.23	[12.49]
Acquirer Leverage			-0.25	[-1.53]			-0.27	[-1.62]
Target Leverage			0.25	[2.05]			0.25	[2.08]
Acquirer M/B			0.01	[0.65]			0.01	[0.79]
Target M/B			0.06	[3.16]			0.06	[3.21]
Target Announcement Return			0.56	[4.40]			0.58	[4.59]
Same Industry			0.08	[1.59]			0.08	[1.52]
Year FE		Yes		Yes		Yes		Yes
N		1,814		1,467		1,814		1,467
R <sup>2</sup>		0.58		0.63		0.57		0.63
Panel B. Cash Deals								
Target IO	0.21	[2.29]	0.11	[0.73]				
Non-Overlapping Target IO					0.27	[1.73]	0.11	[0.52]
Normal Volume	0.99	[9.40]	0.88	[7.92]	0.99	[9.32]	0.88	[7.91]
Additional Controls		No		Yes		No		Yes
Year FE		Yes		Yes		Yes		Yes
N		1,123		899		1,123		899
R <sup>2</sup>		0.44		0.42		0.44		0.42



**Table 7. Acquirer announcement returns: Interactions with proxies for demand-curve slope.** Regressions of acquirer merger announcement returns (in stock deals only) on target institutional ownership, acquirer demand curve proxies, interactions between the two, and deal characteristics. The acquirer announcement return is the return in excess of the value-weighted market over a five-day window surrounding the announcement of the merger. In Panel A, target institutional ownership is totaled from the CDA Spectrum Institutional Holdings database for the quarter prior to the announcement of the merger and expressed as a percentage of shares outstanding. In Panel B, non-overlapping target IO includes only those institutions that own no shares of the acquiring firm and is also expressed as a percentage of shares outstanding. The first two regressions in each panel use dispersion in analyst long-term growth forecasts as a proxy for the slope of the acquirer demand curve; the second two regressions use idiosyncratic risk. The dispersion in analyst forecasts is the standard deviation of all outstanding long-term growth forecasts. Idiosyncratic risk is the standard deviation of the residuals from a regression of acquirer excess returns on the Fama-French Benchmark Factors (Rm, SMB, HML), and the matched 48 industry portfolio return. All factor and portfolio returns were obtained from Ken French's website. The other deal characteristics are described in Table 1. All variables are Winsorized at the first and 99<sup>th</sup> percentiles. The demand curve variables are standardized to have zero mean and unit variance. Announcement year fixed effects are included in all four specifications. Heteroskedasticity-robust t-statistics are reported in braces.

	Dispersion in Analyst Long-Term Growth Forecasts				Non-Industry Idiosyncratic Risk			
	Coef	[t]	Coef	[t]	Coef	[t]	Coef	[t]
<b>Panel A. Target IO</b>								
Target IO	-2.97	[-3.21]	-1.93	[-1.39]	-3.48	[-4.47]	-2.20	[-1.87]
Demand Curve (DC)	0.45	[0.86]	0.67	[1.22]	0.48	[1.39]	0.95	[2.20]
Target IO * DC	-2.62	[-1.85]	-2.62	[-1.78]	-3.41	[-3.02]	-3.22	[-2.73]
Additional Controls		No		Yes		No		Yes
Year FE		Yes		Yes		Yes		Yes
N		1,192		1,001		1,772		1,457
R <sup>2</sup>		0.03		0.09		0.04		0.08
<b>Panel B. Non-Overlapping Target IO</b>								
Non-Overlapping Target IO	-7.19	[-4.12]	-4.72	[-2.29]	-6.69	[-4.80]	-4.71	[-2.75]
Demand Curve (DC)	0.25	[0.50]	0.46	[0.90]	0.19	[0.58]	0.61	[1.44]
Non-Overlapping Target IO * DC	-4.84	[-1.63]	-5.04	[-1.62]	-3.00	[-1.48]	-2.90	[-1.41]
Additional Controls		No		Yes		No		Yes
Year FE		Yes		Yes		Yes		Yes
N		1,192		1,001		1,772		1,457
R <sup>2</sup>		0.04		0.09		0.03		0.08

**Table 8. Financing choice.** Probit regressions of the form of payment (cash or stock) on target institutional ownership, and deal characteristics. Target institutional ownership (IO) is totaled from the CDA Spectrum Institutional Holdings database for the quarter prior to the announcement of the merger and expressed as a percentage of shares outstanding. Non-overlapping target IO includes only those institutions that own no shares of the acquiring firm and is also expressed as a percentage of shares outstanding. The other deal characteristics are described in Table 1. The coefficients show the impact of a unit change in the independent variable, evaluated at its mean, on the probability that the consideration will be the acquiring firm's stock. All variables are Winsorized at the first and 99<sup>th</sup> percentiles. Announcement year fixed effects are included in all four specifications. Heteroskedasticity-robust t-statistics are reported in braces.

	Consideration = Stock							
	Coef	[t]	Coef	[t]	Coef	[t]	Coef	[t]
Target IO	-7.37	[-1.78]	-21.39	[-3.39]				
Non-Overlapping Target IO					-45.61	[-6.77]	-62.51	[-7.27]
log(Acquirer Size)			-3.06	[-4.33]			-4.31	[-5.83]
log(Target Size)			6.34	[6.18]			7.01	[7.68]
Acquirer Leverage			-11.11	[-1.59]			-9.77	[-1.40]
Target Leverage			-5.66	[-1.01]			-6.37	[-1.13]
Acquirer M/B			4.84	[4.71]			4.94	[4.78]
Target M/B			0.91	[1.08]			0.86	[0.99]
Target Announcement Return			-0.43	[-8.17]			-0.41	[-7.88]
Same Industry			13.88	[6.01]			12.59	[5.42]
Year FE		Yes		Yes		Yes		Yes
N		3,003		2,380		3,003		2,380
R <sup>2</sup>		0.04		0.15		0.05		0.16